

Evaluation of Sampling Methods in Constructing Response Surface Approximations

L. P. Swiler^{*}, R. Slepoy, A. A. Giunta

Sandia National Laboratories[†], Albuquerque, NM 87185-0828 USA

Response surface approximations (RSA) are often used as inexpensive replacements for computationally expensive computer simulations. Once a RSA has been computed, it is cheap to evaluate this “meta-model” or surrogate, and thus the RSA is often used in a variety of contexts, including optimization and uncertainty quantification. Usually, some method of sampling points over the input domain is used to generate samples of the input variables. These samples are run through the computer simulation. A response surface approximation is then generated based on the sample points. This report presents a study investigating the dependency of the response surface method on the sampling type. The purpose of this study was to address the question: Does a particular RSA type perform better (in terms of a better fit) if a particular sampling method is used? The RSA types examined were kriging, polynomial regression, and multivariate adaptive regression splines (MARS). The sampling types examined were Latin Hypercube, Halton, Hammersley, Centroidal Voronoi Tessellation (CVT), and standard Monte Carlo sampling. The example problems were a 5-dimensional version of Rosenbrock’s function and the Paviani function. RSA of the three response surface types were developed based on the five sampling methods. Performance was compared using ANOVA techniques.

^{*} Corresponding Author: US Postal Address: Sandia National Laboratories, P.O. Box 5800, Mail Stop 0370, Albuquerque, NM 87185-0370, USA; Email: lpswire@sandia.gov; Telephone: 505-844-8093

[†] Sandia is a multiprogram laboratory operated by Sandia Corporation, a Lockheed Martin Company, for the United States Department of Energy’s National Nuclear Security Administration under Contract DE-AC04-94AL85000.

This material is declared a work of the U.S. Government and is not subject to copyright protection in the United States.

I. Introduction

Response surface approximations (RSA) are often used as inexpensive replacements for computationally expensive computer simulations. Once a RSA has been computed, it is cheap to evaluate this “meta-model” or surrogate, and thus the RSA is often used in a variety of contexts, including optimization and uncertainty quantification. Usually, some method of sampling points over the input domain is used to generate samples of the input variables. These samples are run through the computer simulation. A response surface approximation is then generated based on the sample points.⁸

This report presents a study investigating the dependency of the response surface method on the sampling type. The purpose of this study was to address the question: Does a given RSA type perform better (in terms of a better fit) if a particular sampling method is used? There is some evidence to suggest that quasi-Monte Carlo methods perform better than Latin Hypercube when sampling over a small number of input variables¹⁷. There is also some evidence that kriging does not perform well when the points are highly collinear¹. Thus, we wanted to investigate these issues in more detail and provide some guidance within the context of the DAKOTA software framework on which sampling methods work best for various RSA types.

The RSA types examined were kriging, polynomial regression, and multivariate adaptive regression splines (MARS). The sampling types examined were Latin Hypercube, Halton, Hammersley, Centroidal Voronoi Tessellation (CVT), and standard Monte Carlo sampling. The example problems were a 5-dimensional version of Rosenbrock’s function and a 10-dimension version of Paviani’s function. RSA of the three response surface types were developed based on the five sampling methods.

II. Sampling Methods

The sampling methods presented in sections 2.1-2.5 below all can be used in uncertainty quantification, where one wants to understand the effect of input uncertainties on the distribution of outputs. As an example, consider a variable Y that is a function of other variables X_1, X_2, \dots, X_k . This function may be very complicated, for example, a computer model. A question to be investigated is “How does Y vary when the X s vary according to some assumed joint probability distribution?” Related questions are “What is the expected value of Y ?” and “What is the 99th percentile of Y ?” The sampling methods can all generate a number of samples of the random variables. It is convenient to think of the samples as forming an $(n \times k)$ matrix of input where the i^{th} row contains specific values of each of the k input variables to be used on the i^{th} run of the computer model.

A. Latin Hypercube

Latin hypercube sampling (LHS) is a stratified sampling method developed to address the need for more efficient uncertainty assessment. LHS partitions the parameter space into bins of equal probability, with the goal of attaining a more even distribution of sample points in the parameter space than typically occurs with pure Monte Carlo sampling.

Latin hypercube sampling, developed by McKay, Conover, and Beckman¹⁵, is a constrained sampling method which selects n different values from each of k variables X_1, \dots, X_k in the following manner. The range of each variable is divided into n non-overlapping intervals on the basis of equal probability. One value from each interval is selected at random with respect to the probability density in the interval. The n values thus obtained for X_1 are paired in a random manner (equally likely combinations) with the n values of X_2 . These n pairs are combined in a random manner with the n values of X_3 to form n triplets, and so on, until n k -tuplets are formed. This is the Latin hypercube sample.

The Latin hypercube sampling technique has been applied to many different computer models over the past thirty years. A tutorial on Latin hypercube sampling may be found in Iman and Conover¹². A recent comparison of Latin hypercube sampling with other techniques is given in Helton and Davis¹¹. A method for inducing correlations among the input variables is given in Iman and Conover¹³. Swiler and Wyss²¹ provide a User’s manual for the current LHS implementation in the DAKOTA software framework.

B. Halton Sampling

Halton sampling is a type of sampling known as quasi-Monte Carlo sampling. The goal of quasi-Monte Carlo methods is to produce sequences which have low discrepancy. Discrepancy refers to the nonuniformity of the sample points within the hypercube. Discrepancy is defined as the difference between the actual number and the expected number of points one would expect in a particular set B (such as a hyper-rectangle within the unit hypercube), maximized over all such sets. Low discrepancy sequences tend to cover the unit hypercube reasonably uniformly. Quasi-Monte Carlo methods produce low discrepancy sequences, especially if one is interested in the uniformity of projections of the point sets onto lower dimensional faces of the hypercube (usually 1-D: how well do the marginal distributions approximate a uniform?)

The quasi-Monte Carlo Halton sequence is a deterministic sequence determined by a set of prime bases. The Halton sequence in base 2 starts with points 0.5, 0.25, 0.75, 0.125, 0.625, etc. The first few points in a Halton base 3 sequence are 0.33333, 0.66667, 0.11111, 0.44444, 0.77777, etc. Notice that the Halton sequence tends to alternate back and forth, generating a point closer to zero then a point closer to one. An individual sequence is based on a radix inverse function defined on a prime base. The prime base determines how quickly the [0,1] interval is filled in. Generally, the lowest primes are recommended.

For more information about the Halton sequence, see References 9, 10, and 14. In cases where a large number of input variables are sampled, Robinson and Atcitty¹⁷ recommend using a leaped sequence, where the user does not use every term in the Halton sequence but sets a “leap value” to the next prime number larger than the largest prime base. Using the leaped values in the sequence can help maintain uniformity when generating sample sets for high dimensions.

C. Hammersley Sampling

The Hammersley sequence is the same as the Halton sequence, except the values for the first random variable are equal to $1/N$, where N is the number of samples. Thus, if one wants to generate a sample set of 100 samples for 3 random variables, the first random variable has values $1/100$, $2/100$, $3/100$, etc. and the second and third variables are generated according to a Halton sequence with bases 2 and 3, respectively. Hammersley sequences can also be improved but shifting points³. Since we are using a relatively low number of dimensions for this study (five), we did not use a shifted or leaped version of either the Halton or Hammersley sequences.

D. Centroidal Voronoi Tessellation

Centroidal Voronoi Tessellation (CVT) sampling produces a set of sample points that are (approximately) a Centroidal Voronoi Tessellation. The primary feature of such a set of points is that they have good volumetric spacing; the points tend to arrange themselves in a pattern of cells that are roughly the same shape. To produce this set of points, an almost arbitrary set of initial points is chosen, and then an internal set of iterations is carried out. These iterations repeatedly replace the current set of sample points by an estimate of the centroids of the corresponding Voronoi subregions.⁴

CVT does very well volumetrically: it spaces the points fairly equally throughout the space, so that the points cover the region and are isotropically distributed with no directional bias in the point placement. There are various measures of volumetric uniformity which take into account the distances between pairs of points, regularity measures, etc. Note that CVT does not produce low-discrepancy sequences in lower dimensions, however: the lower-dimension (such as 1-D) projections of CVT can have high discrepancy.¹⁸

E. Monte Carlo

What we are referring to as “pure” or “plain” Monte Carlo¹⁶ involves using a random number generator to generate random number sequences, with no effort to stratify the samples or construct explicit correlations between sample values for multiple input dimensions.

F. DAKOTA Implementation

This study employs DAKOTA⁵ version 3.3 on a 32-bit Intel microprocessor-based computer workstation running the Fedora Core 3 version of the Red Hat Linux operating system. This version of DAKOTA is available to the public, under the restrictions of the GNU General Public License, from <http://endo.sandia.gov/DAKOTA>.

III. Response Surface Approximation Methods

A. Kriging

Kriging interpolation techniques were originally developed in the geostatistics and spatial statistics communities to produce maps of underground geologic deposits based on samples obtained at widely and irregularly spaced borehole sites². The basic notion that underpins kriging is that the sample response values exhibit spatial correlation, with response values modeled via a Gaussian process around each sample location (i.e., samples taken close together are likely to have highly correlated response values, whereas samples taken far apart are unlikely to have highly correlated response values). Kriging methods have found wide utility due to their ability to accommodate irregularly spaced data, their ability to model general surfaces that have many peaks and valleys, and their exact interpolation of the given sample response values.

The specific form of the kriging model used in this study is described in Giunta and Watson⁷ and Romero et al.¹⁹ The form of the kriging model is

$$\hat{f}(x) = \beta + r(x)^T R^{-1}(f - \beta\{1\}), \quad (1)$$

where β is the generalized least squares estimate of the mean response; $r(x)$ is an $N \times 1$ vector of correlations between the current point x , and all N sample sites in parameter space; R is the $N \times N$ correlation matrix of all N sample sites; f is the vector of N sample site response values; and $\{1\}$ is an $N \times 1$ vector with all values set to unity. The terms in the correlation vector and matrix are computed using a Gaussian correlation function. The i^{th} term in $r(x)$ is given by

$$r_i = \exp\left[-\sum_{t=1}^n \theta_t |x_t - x_t^{(i)}|^2\right], \quad (2)$$

and, similarly, the i,j^{th} term in R is given by

$$R_{i,j} = \exp\left[-\sum_{t=1}^n \theta_t |x_t^{(i)} - x_t^{(j)}|^2\right], \quad (3)$$

where n is the dimension of the parameter space; t is the index on the dimension of the parameter space; $i = 1, \dots, N$; $j = 1, \dots, N$; and θ is the $n \times 1$ vector of correlation parameters. In this study, all values of θ are set to unity, although in general, the values of θ can be estimated from the N sample response values via maximum likelihood estimation.

However, there are drawbacks to kriging. The form of the kriging model requires the inversion of a potentially dense $N \times N$ matrix, where N is the number of sample points. Thus, the basic kriging method does not scale well for large N . In addition, if two or more of the sample points are close together, the $N \times N$ matrix becomes ill-conditioned. Thus, while kriging tends to work well for sparse sets of samples, this method tends to break down as the number of samples increases. Note that the kriging interpolation method is prone to ill-conditioning in the correlation matrix R as the number of sample points increases. This occurs because of the distance measure that is computed in Equation (3). As the distance between any two sample points i and j decreases, then the i^{th} and j^{th} rows in matrix R become linearly dependent, and in the limit where the points are the same, the matrix R becomes singular. Thus, this basic kriging method works well for a sparse set of sample points in an n -dimensional parameter space, but as the number of samples increases (and the inter-point distances decrease), the kriging method becomes unstable.

B. Polynomial Regression

Polynomial regression methods are commonly used to create RSA from a set of data samples. Regression is popular since the calculations are simple and the resulting function is a closed-form algebraic expression. For example, a quadratic polynomial has the form:

$$\hat{Y}(X) = C_0 + \sum_{i=1}^k C_i X_i + \sum_{i=1}^k \sum_{j=1}^k C_{i,j} X_i X_j \quad (4)$$

Where $\hat{Y}(X)$ is the estimate of target function at X and the C_0 , C_i , $C_{i,j}$, are constant coefficients. To calculate the values of the coefficients, a system of linear equations is formed by applying the above polynomial model at each of N sampling points. The number of sampling points must be greater than or equal to the number of

ascending terms in the polynomial. If equal, then the system of equations is exactly determined and the coefficients of the saturated polynomial can be immediately solved for and the resulting RSA exactly matches the target values at all of the sampling points. When more sampling points than terms in the polynomial exist, then the system of equations is overdetermined, and a regression procedure is invoked to solve for the coefficients. Because the regression polynomial does not have as many terms as there are data samples to fit, it is over-constrained and cannot in general match the true Y values at the sample points. Most commonly, the method of least squares is used.

C. Multivariate Adaptive Regression Splines

The multivariate adaptive regression splines (MARS) function approximation method⁶ is based on a complex, recursive partitioning algorithm involving truncated power spline basis functions. The form of the MARS model is:

$$\hat{f}(x) = a_o + \sum_{m=1}^{M_1} a_m B_m(x_i) + \sum_{m=1}^{M_2} a_m B_m(x_i, x_j) + \dots \quad (5)$$

where the B_m terms are the basis functions, the a_m terms are the coefficients of the basis functions, M_1 is the number of one-parameter basis functions, and M_2 is the number of two-parameter basis functions. The MARS software allows the user to select either linear or cubic spline basis functions. Cubic spline basis functions are used for this study. The regression aspect of the MARS algorithm involves a forward/backward stepping process to adaptively add/remove spline basis functions from the model. It is this regression process that generates the a_o and a_m terms in Equation (5). The resulting MARS model is a C^2 -continuous function of piecewise cubic splines, but it will not exactly interpolate the data points that were used in calculating the coefficients. Thus, like polynomial regression, MARS has the ability to create smooth approximations to noisy data. Unlike kriging, MARS appears to have no upper limit on the number of samples that can be used in the function approximation process.

IV. Results

A. Analysis Approach

We used Analysis of Variance (ANOVA) methods to determine if the means of various statistics of interest (such as root mean square error) were significantly different when using one sampling method vs. another, for a given response surface type. Analysis of variance is used to examine the correlation between a response variable (in our case, measures of response surface goodness-of-fit) and the independent variables (in our case, sampling method). ANOVA extends the two-sample t-test for testing the equality of two population means to a more general null hypothesis of comparing the equality of more than two means.

We used two goodness-of-fit measures to evaluate the accuracy of a response surface constructed from a particular set of sample values: root mean squared error (RMSE) and mean absolute error (MAE). The definition of these terms is given below, where y_i = actual or observed value and \hat{y}_i is the value predicted by the response surface.

$$\text{RMSE} = \sqrt{\frac{1}{n} \sum_{i=1}^n (y_i - \hat{y}_i)^2}$$

$$\text{MAE} = \frac{1}{n} \sum_{i=1}^n \left| \frac{y_i - \hat{y}_i}{y_i} \right|$$

Note that to evaluate each of these sampling methods, we generated 50 sample sets. Each particular sample set had 100 sample points in 5 dimensions. We evaluated a 5-dimensional version of the Rosenbrock function at these 100 points. The function is:

$$f(x) = \sum_{i=1}^4 [100(x_{i+1} - x_i^2)^2 + (x_i - 1)^2]$$

After evaluating the function at the 100 points, we constructed a response surface based on these points, then used the response surface to predict the values of the Rosenbrock function on a grid. The grid had 9 “levels” of each

of the 5 input variables, so the grid had 59049 points. We then constructed error metrics for each of the 50 sample sets (e.g., each sample set has one value for RMSE and MAE). The ANOVA tests are looking across the 50 samples, to see if the mean RMSE according to LHS sampling (for example) is different from the mean RMSE according to Halton sampling, etc.

To illustrate the process, each sampling method was used to generate fifty sample sets, each with 100 samples in 5-d. For example, one of the LHS input sets was:

X1	X2	X3	X4	X5	Rosenbrock Fn.
-0.849	1.593	-0.983	-1.417	-0.676	2617.739
-1.752	1.978	-0.293	0.710	0.591	1937.145
0.987	-1.172	-1.301	-0.848	1.945	1986.069
-0.476	-1.833	-1.623	-1.133	1.252	4346.873
-1.522	-0.217	1.781	1.698	0.693	1649.052
-0.425	-1.216	1.296	-1.083	-0.316	1194.997
0.590	0.839	-1.261	-1.732	-0.873	3024.397
0.653	1.712	-1.990	-1.452	0.936	5672.056
1.277	-1.284	0.514	-0.188	-1.565	1262.108
-1.347	-0.844	-0.816	0.461	0.326	957.117

...

Note that the input variables are all bounded between -2 and 2. The predictions were constructed on a 5-d grid as follows (this example shows a prediction based on a kriging response surface constructed over the input samples):

X1	X2	X3	X4	X5	Rosenbrock Fn.	Kriging Prediction
-2.00	1.50	2.00	2.00	1.50	1667.50	1186.91
-1.50	1.50	2.00	2.00	1.50	1096.00	1186.22
-1.00	1.50	2.00	2.00	1.50	1062.50	1185.58
-0.50	1.50	2.00	2.00	1.50	1192.00	1184.99
0.00	1.50	2.00	2.00	1.50	1259.50	1184.46
0.50	1.50	2.00	2.00	1.50	1190.00	1183.99
1.00	1.50	2.00	2.00	1.50	1058.50	1183.58
1.50	1.50	2.00	2.00	1.50	1090.00	1183.22
2.00	1.50	2.00	2.00	1.50	1659.50	1182.92
-2.00	2.00	2.00	2.00	1.50	1837.00	1241.43
-1.50	2.00	2.00	2.00	1.50	1440.50	1241.05
-1.00	2.00	2.00	2.00	1.50	1532.00	1240.69
-0.50	2.00	2.00	2.00	1.50	1736.50	1240.37
0.00	2.00	2.00	2.00	1.50	1829.00	1240.07
-2.00	1.50	2.00	2.00	1.50	1667.50	1186.91

The difference between the actual Rosenbrock function value and the predicted function value was used in to construct an RMSE and MAE metric for each sample set. Then, we used ANOVA to understand the spread of these metrics over the 50 sample sets, and to determine if the means were the same. Below we present box plots and ANOVA results for each response surface type.

B. Results for Kriging RSA

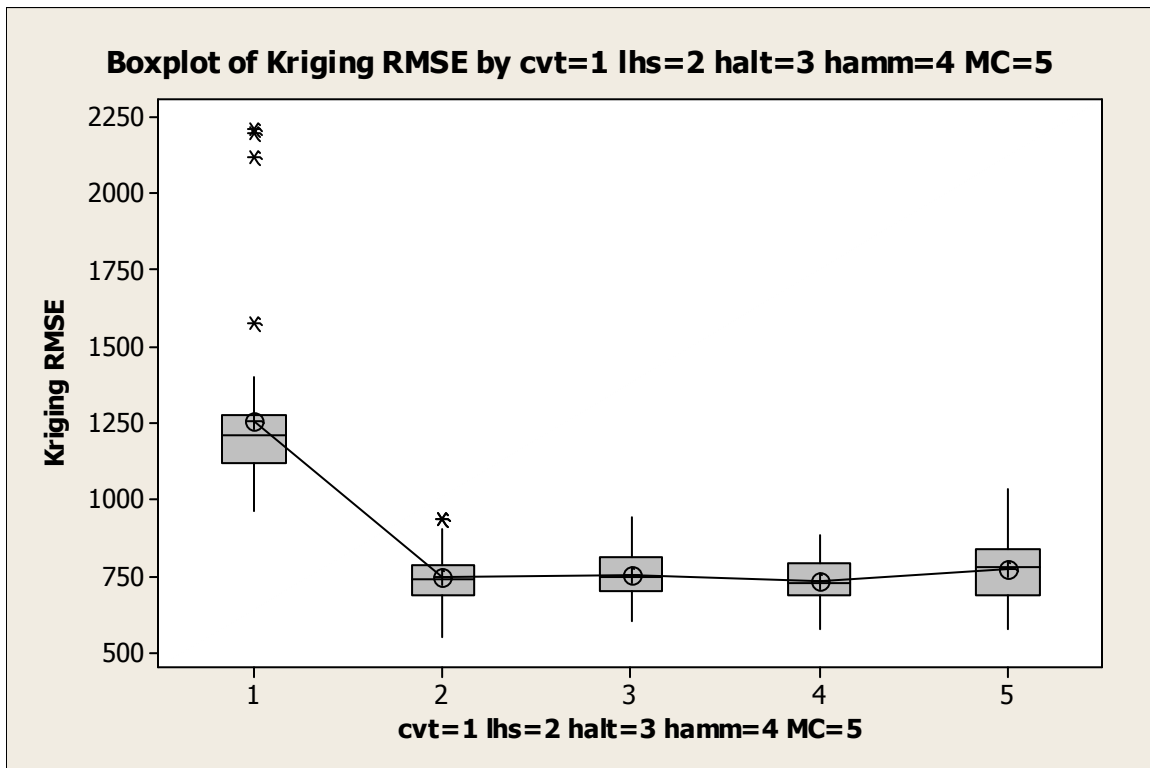


Figure 1. Kriging RMSE vs. Sample Type: Rosenbrock Function

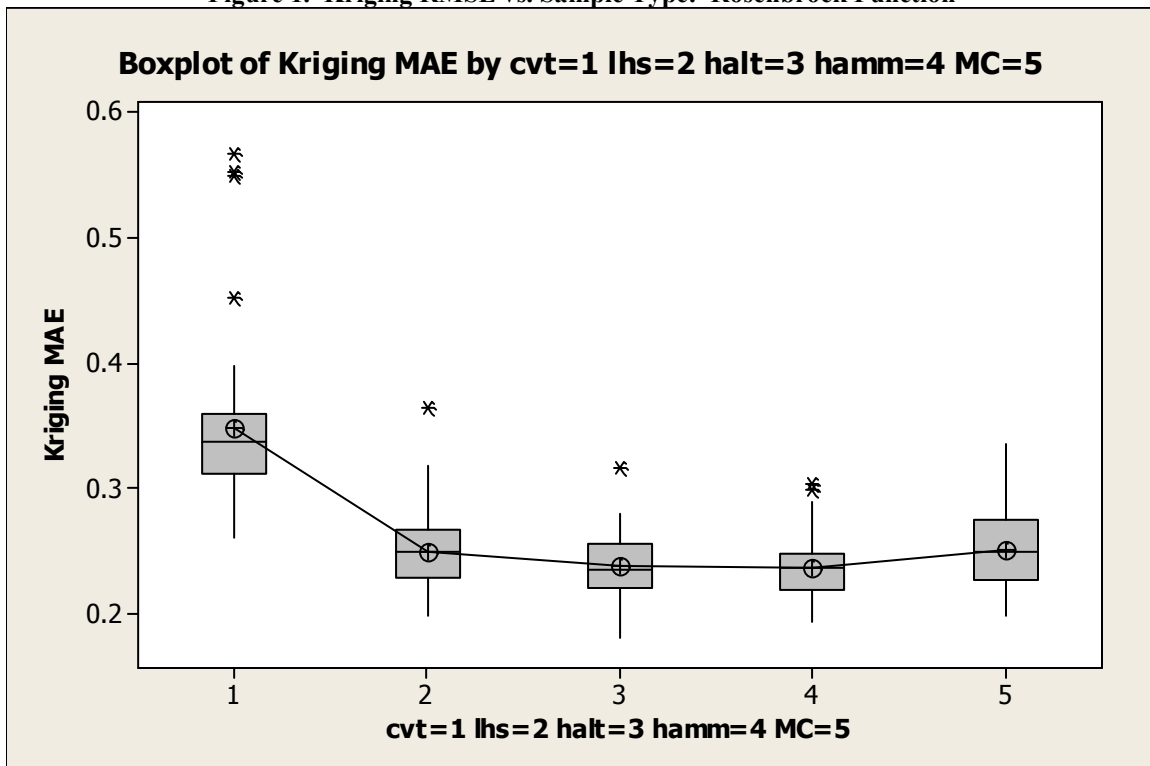


Figure 2. Kriging MAE vs. Sample Type: Rosenbrock Function

Level	N	Mean	StDev	Individual 95% CIs For Mean Based on Pooled StDev
1	50	1251.9	260.8	(- * - -)
2	50	747.1	88.2	(- * -)
3	50	748.4	77.5	(- * -)
4	50	730.0	77.7	(- * -)
5	50	773.1	109.7	(- * - -)

59K Residuals from a Kriging model
 based on various sample types: CVT, LHS, Halton, Hammersley, and MC

CVT

LHS

Halton

Hammersley

MC

-2500 0 2500 5000 7500 10000 12500

Data

Each symbol represents up to 2614 observations.

8
American Institute of Aeronautics and Astronautics

C. Results for Polynomial Regression RSA

Figure 4 shows a similar pattern for the various sample types when a polynomial regression response surface is constructed based on these sample types. We see that the mean value of RMSE is statistically significantly greater for CVT than for the other sampling methods when using a polynomial regression response surface.

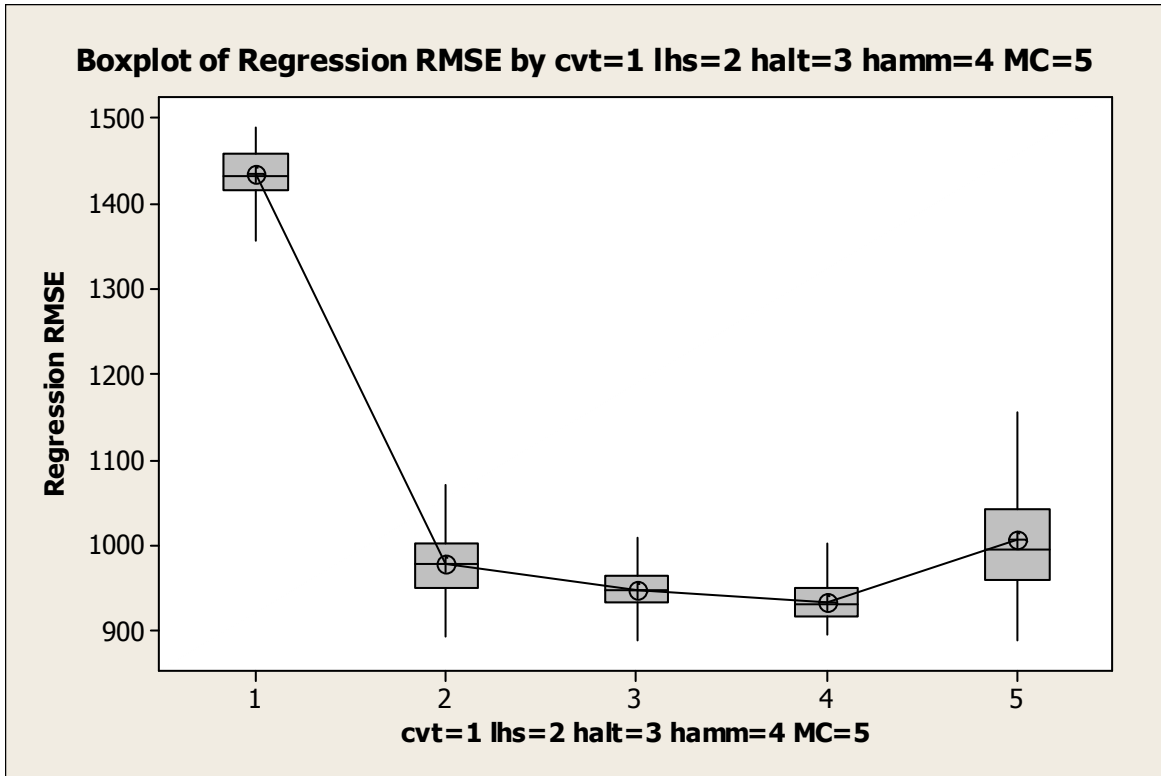
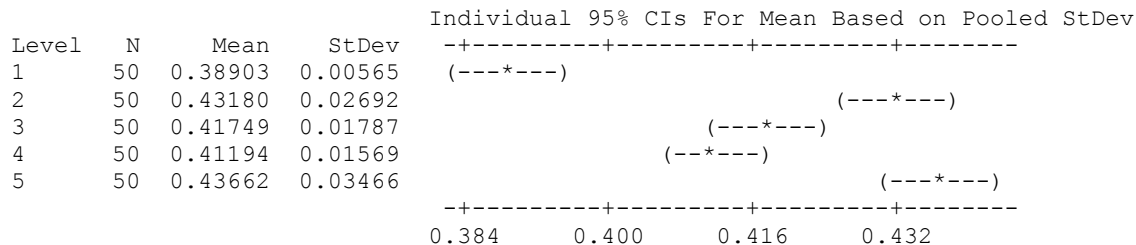


Figure 4. Polynomial Regression RMSE vs. Sample Type: 5-D Rosenbrock

However, the MAE for CVT is not greater than the other sampling methods when a polynomial regression model is used, in fact the reverse is true, as shown in Figure 5. The MAE for CVT is the smallest of all the sampling methods and the mean MAE based on these 50 samples is statistically significantly smaller than the mean MAEs from the other sampling methods. Also, Hammersley and Halton performed somewhat better than LHS and MC for MAE:



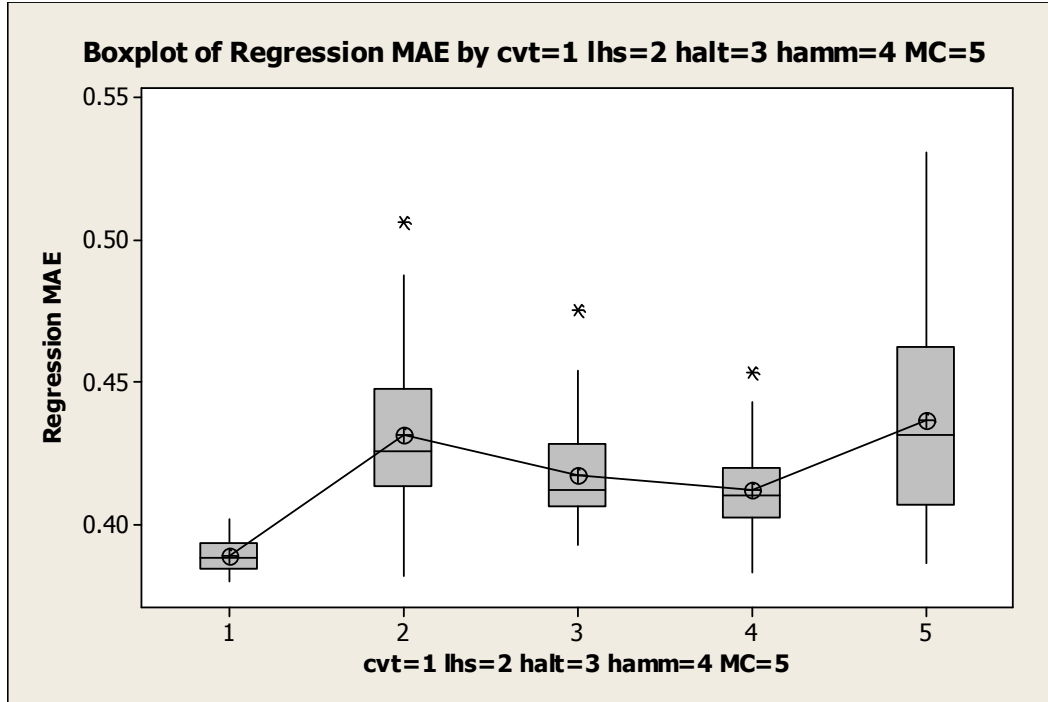


Figure 5. Polynomial Regression MAE vs. Sample Type: 5-D Rosenbrock

Although the “good” performance of CVT with respect to the mean absolute error metric may seem contradictory to the “poor” performance of CVT with respect to the root mean squared error metric, one needs to remember that these metrics track different behavior. The Rosenbrock function is fairly flat in the middle of the domain chosen for this study, but its value increases sharply at the edges of the domain. All of the response surface methods did poorly at predicting the 5-D Rosenbrock function at the edges of the domain. Significant errors in “edge” fitting can lead to large RMSE, while average error is not affected much by lack of fit near the edges (remember that we are taking average error over 59049 points). Also, it is instructive to look at the placement of the sample points upon which the regression surface is based. If one looks at a plot of the CVT points in two dimensions, for example inputs X_1 and X_2 , one can see they are “clustered” as shown in Figure 6. Plotting any of the other inputs relative to each other (e.g., X_3 vs. X_5) shows a similar pattern. This clustering may contribute to the method performing relatively well over all the space but poorly at the edges, which the RMSE metric emphasizes. Note that there is an approach which “latinizes” or stratifies the CVT samples to give them better 1-D marginal densities, which may improve their potential use in response surface modeling. This is described in Reference 18 and is implemented in the DAKOTA software. However, we did not latinize the CVT samples for the purposes of this study.

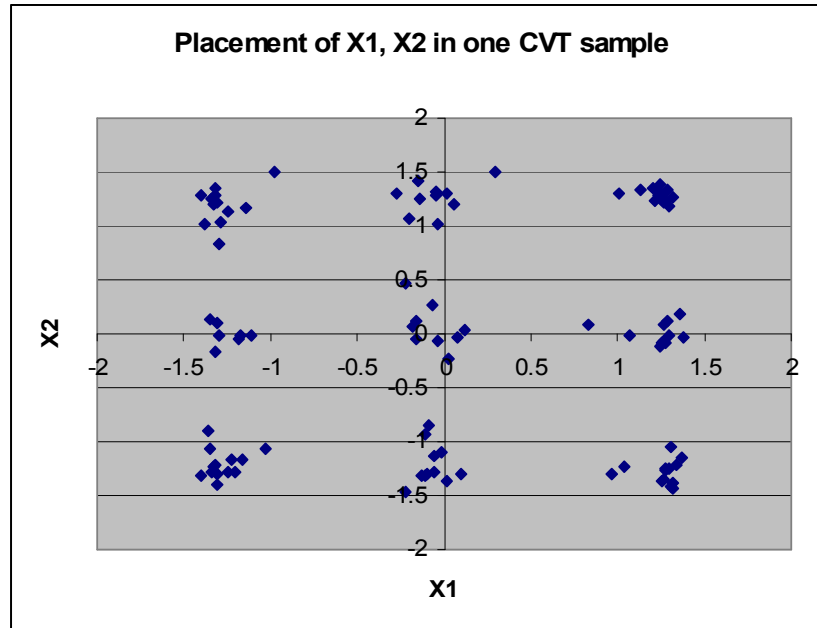


Figure 6. Placement of X1, X2 in a 100 point CVT sample set

D. Results for Multivariate Adaptive Regression Spline RSA

The MARS results in Figure 7 again show that CVT performed poorly relative to the other sampling methods when comparing RMSE. The performance of the other sampling methods showed minor differences, with Monte Carlo performing worse than LHS, Hammersley, and Halton.

Level	N	Mean	StDev	Individual 95% CIs For Mean Based on Pooled StDev	
1	50	1541.7	101.6	-----+-----+-----+-----	(-*)
2	50	826.4	144.7	(*-)	
3	50	806.8	154.7	(*-)	
4	50	756.8	121.6	(*-)	
5	50	891.7	158.9	(-*)	
				-----+-----+-----+-----	
				750 1000 1250 1500	

With MARS MAE, CVT did perform worse than the other sampling methods as shown by the ANOVA results below and Figure 8:

Level	N	Mean	StDev	Individual 95% CIs For Mean Based on Pooled StDev	
1	50	0.46926	0.03771	-----+-----+-----+-----	(--*--)
2	50	0.34834	0.06711	(--*--)	
3	50	0.33059	0.06189	(--*--)	
4	50	0.31438	0.04752	(--*--)	
5	50	0.35646	0.05639	(--*--)	
				-----+-----+-----+-----	
				0.300 0.350 0.400 0.450	

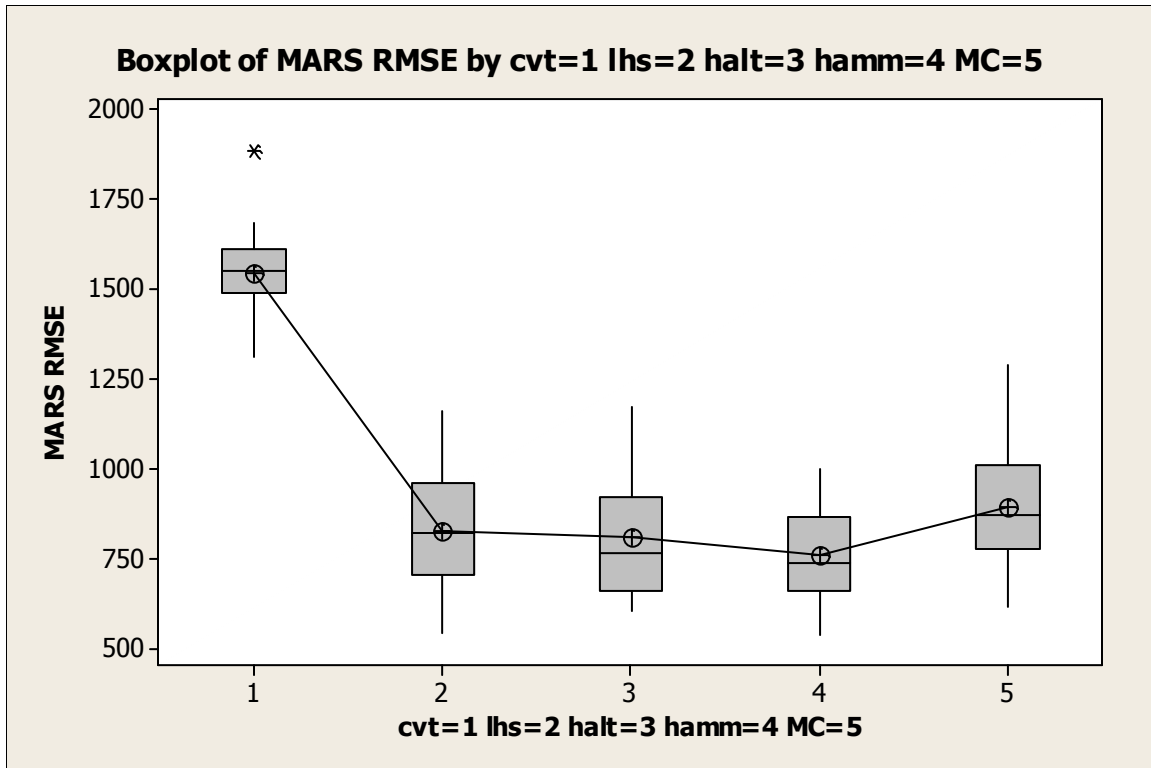


Figure 7. MARS RMSE vs. Sample Type: 5-D Rosenbrock

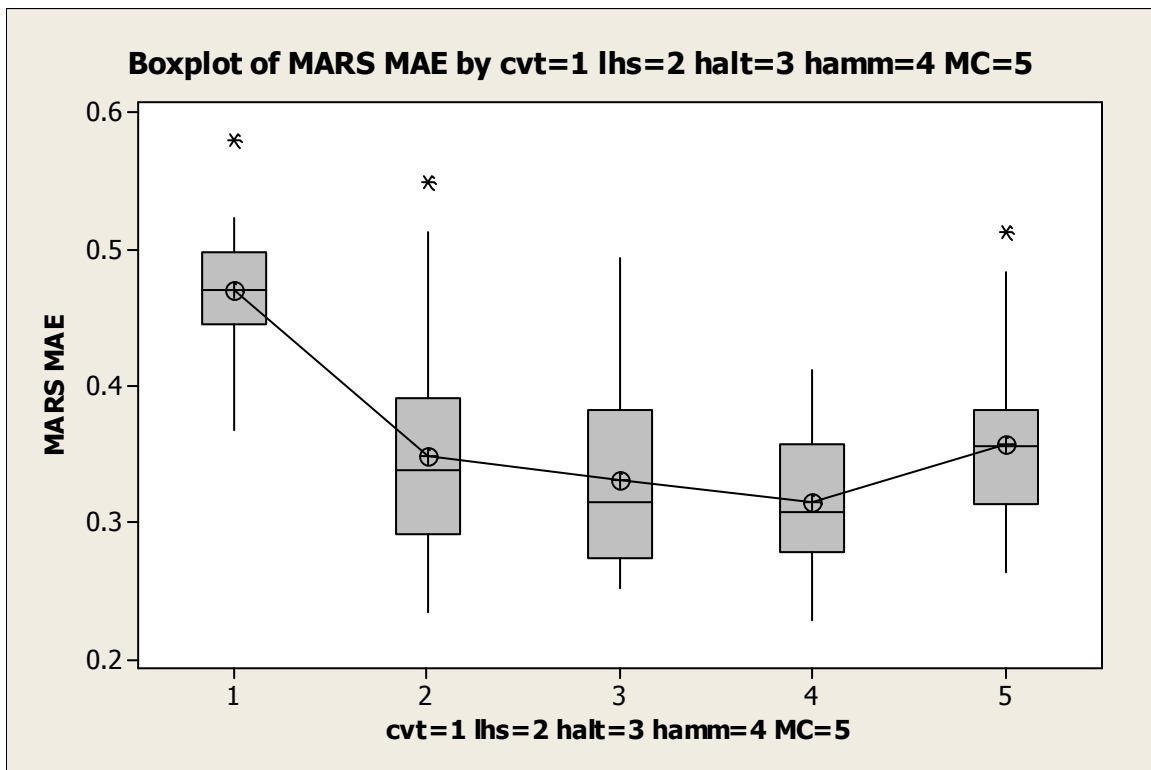


Figure 8. MARS MAE vs. Sample Type: 5-D Rosenbrock

Overall, kriging and MARS performed better than regression on the 5-D Rosenbrock function across all of the sample types except CVT. For CVT, kriging and regression performed better than MARS. Figure 9 shows a 2-way ANOVA, with the details below.

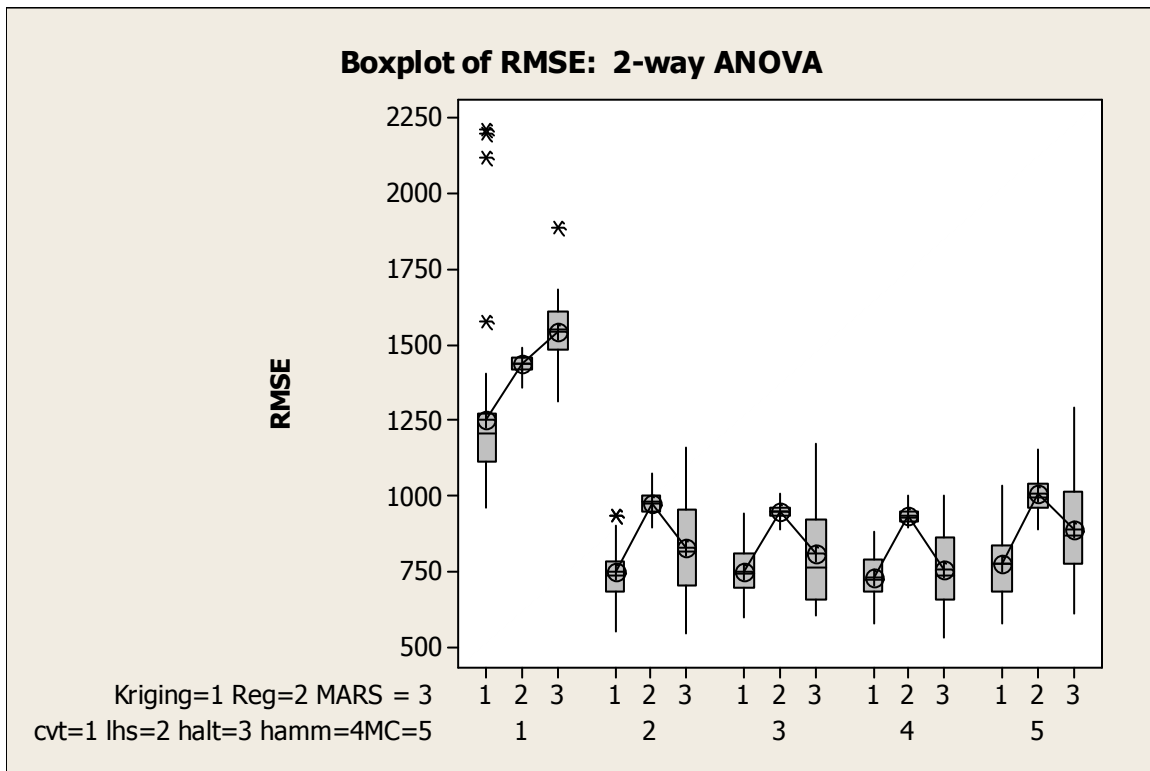


Figure 9. Two-way ANOVA for 5-D Rosenbrock Function

Two-way ANOVA: RMSE versus cvt=1 lhs=2 halt = 3 hamm=4 MC=5, Kriging=1 Reg=2 MARS =3

Source	DF	SS	MS	F	P
cvt=1 lhs=2 halt	4	38675254	9668814	717.25	0.000
Kriging=1 Reg=2	2	5527618	2763809	205.02	0.000
Interaction	8	1636826	204603	15.18	0.000
Error	735	9908081	13480		
Total	749	55747779			

The two-way ANOVA table tells us that the effect of the differences in the sample type has the greatest effect on the variance of RMSE, followed by the response surface type. There is an interaction effect when CVT is included in the sample types. However, if we remove CVT, then the analysis changes, as shown below. With the removal of CVT, there is not a significant interaction between sample type and response surface method at the 0.05 level of significance. This means that the performance of the response surface methods at a different level (regression, MARS, kriging) do not depend on the sampling types. Also, the response surface type has become the most important at affecting the variance of RMSE. This is shown by the SS (sum of squares) which measures how much the level means differ within each factor.

**Two-way ANOVA: RMSE versus lhs=2 halt=3 hamm=4 MC=5, Kriging=1 Reg=2
MARS = 3**

Source	DF	SS	MS	F	P
lhs=2 halt=3 ham	3	549074	183025	17.83	0.000
Kriging=1 Reg=2	2	4895728	2447864	238.44	0.000
Interaction	6	122194	20366	1.98	0.066
Error	588	6036389	10266		
Total	599	11603386			

S = 101.3 R-Sq = 47.98% R-Sq(adj) = 47.00%

Two-way analysis of MAE leads to the same result when CVT is removed: there is no significant interaction between sampling type and response surface type:

**Two-way ANOVA: MAE versus lhs=2 halt=3 hamm=4MC=5, Kriging=1 Reg=2
MARS = 3**

Source	DF	SS	MS	F	P
lhs=2 halt=3 ham	3	0.07291	0.02430	14.93	0.000
Kriging=1 Reg=2	2	3.28768	1.64384	1010.19	0.000
Interaction	6	0.01010	0.00168	1.03	0.402
Error	588	0.95683	0.00163		
Total	599	4.32751			

S = 0.04034 R-Sq = 77.89% R-Sq(adj) = 77.48%

An important point to note with the 5-D Rosenbrock function is that none of the response surface methods performed very well with respect to either RMSE or MAE. The values of this function range from 0 to 14436 over the input domain $[-2,2]^5$. The histogram is shown in Figure 10.

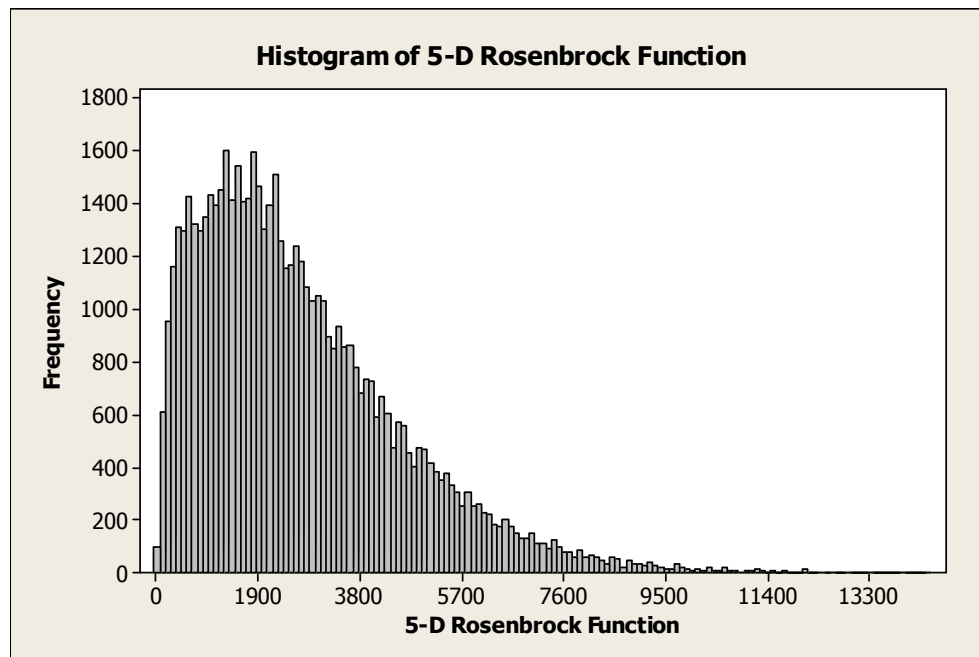


Figure 10. Histogram of 5-D Rosenbrock Function

Because the function values vary so widely over a fairly small input domain, these response surface methods are not very accurate. A root mean square error of 700, for example, or a mean absolute error of 40% which were demonstrated across sampling methods and response surface approximation types, shows the inaccuracy of these

response surfaces. We chose this function to mimic a nonlinear response and to provide an extreme test case. The next function we chose, the Paviani function, is better behaved in terms of the output not varying widely over the input domain.

E. The Paviani Function

After we performed the analysis on the Rosenbrock function, we wanted to perform a similar analysis on a higher dimensional function with different characteristics. As mentioned, the Rosenbrock function is difficult to approximate with response surface methods as the function value sharply increases near the corners of the region $[-2,2]^5$. We chose the Paviani function because it is a nonlinear function which has interactions between all 10 input variables. The Paviani function in 10 dimensions is:

$$f(x) = \sum_{i=1}^{10} (\ln^2(x_i - 2) + \ln^2(10 - x_i)) - \left(\prod_{i=1}^{10} x_i\right)^{0.2}.$$

The input domain is defined as $[2.001, 9.999]^{10}$. The minimum of this function is -45.7785 at the point where all of the input values equal 9.3503. Figures 11 and 12 show scatterplots of x_1, x_2 vs. the Paviani function and x_5, x_6 vs. the Paviani function.

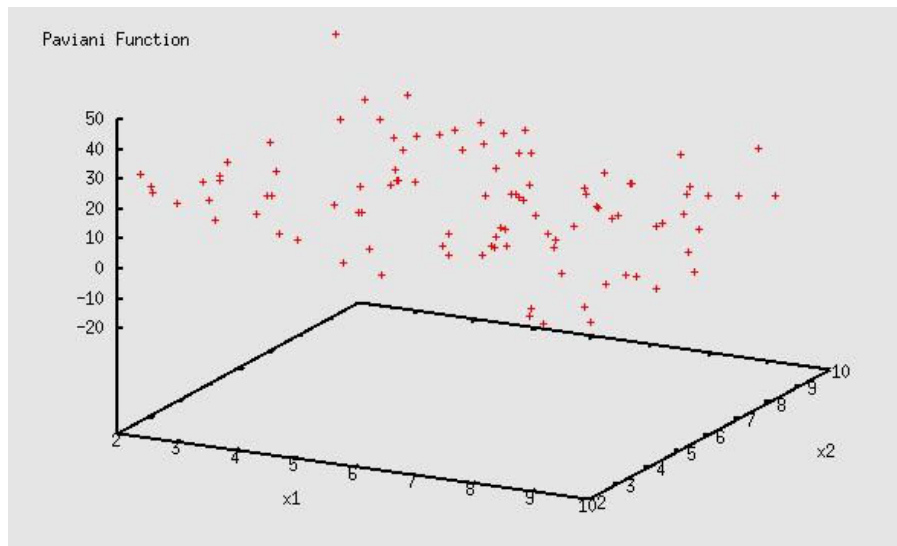


Figure 11. X1 vs. X2 vs. the Paviani Function

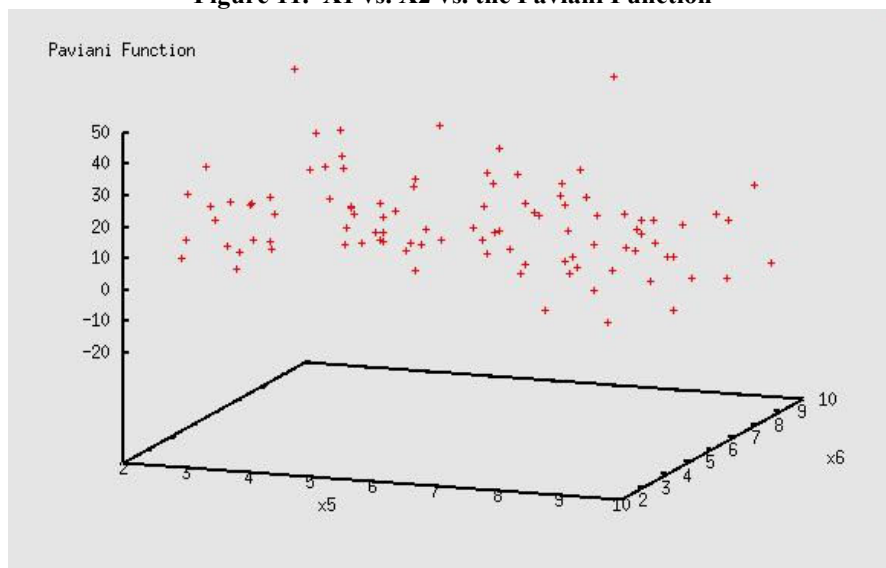


Figure 12. X5 vs. X6 vs. the Paviani Function

We used a similar procedure as the Rosenbrock function for performing the analysis: We generated 50 sample sets, where each sample set had 100 samples in the 10-dimensional input space. With each of the 100 samples, we evaluated the Paviani function, and used those 100 values to construct kriging, regression, and adaptive spline response surfaces. Then we used the response surfaces to evaluate the function at a grid. The grid was constructed at three levels for each of the 10 inputs: 3, 5.5, and 8. There were 59049 gridpoints which were used to construct the metrics such as RMSE and MAE. Note that for the Paviani study, we chose some of the gridpoints to be slightly on the interior of the domain, whereas for the Rosenbrock function, some of the gridpoints were at the outer boundaries of the domain.

The results are shown in Sections 4.6-4.8 below. The most striking result is the reversal of CVT. In the Rosenbrock studies, CVT did not perform well as a sampling method, especially in terms of the RMSE metric. In the Paviani studies, CVT often outperformed the other sampling methods significantly.

F. Results of Kriging RSA on the Paviani function

Figures 13 and 14 below shows the results of the RMSE and the MAE for the kriging RSA of the Paviani function. The important thing to note is that CVT does not perform well with respect to RMSE, but it performs very well if the goal is to minimize the mean absolute error of the response surface over the grid. The CVT results are statistically significantly different than the rest of the sampling methods both for RMSE and for MAE.

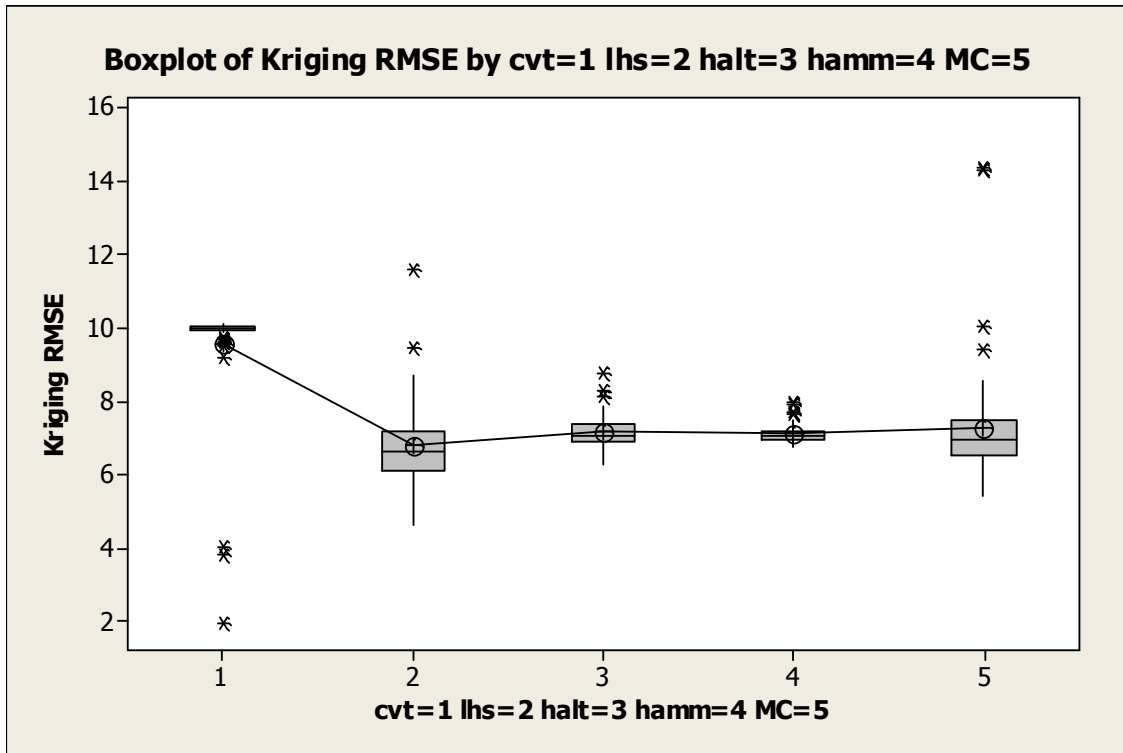


Figure 13. Kriging RMSE vs. Sample Type: Paviani function

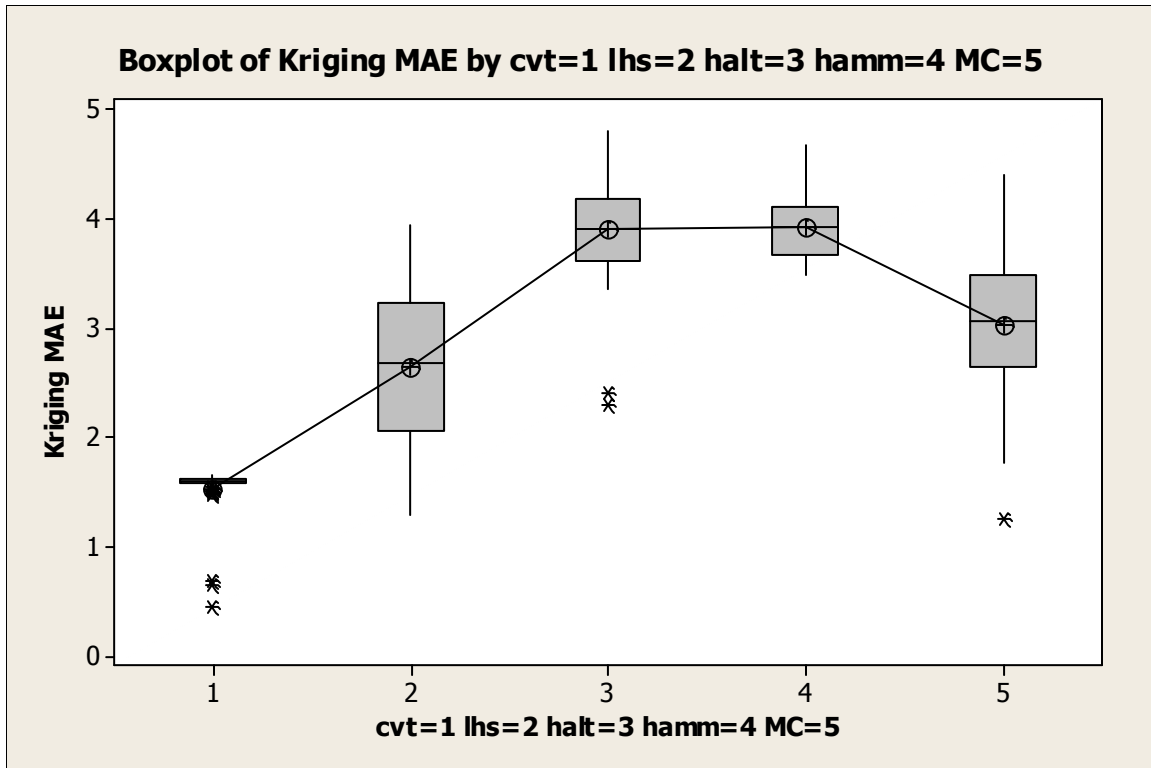


Figure 14. Kriging MAE vs. Sample Type: Paviani function

G. Results of Polynomial Regression RSA on the Paviani function

Figure 15 shows the results of the RMSE for the kriging RSA of the Paviani function. The Hammersley function performed horribly, with several values of RMSE on the order of 10^9 .

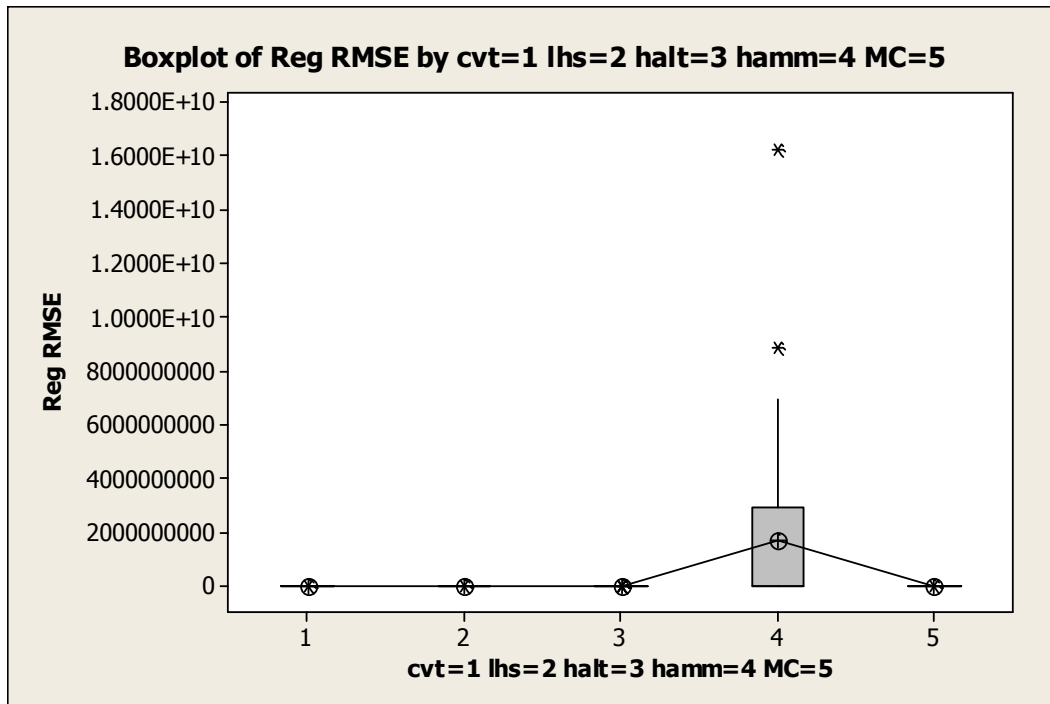


Figure 15. Polynomial Regression RMSE vs. Sample Type: Paviani function

The likely reason for this is that all 10 variables are important in the Paviani function. In Hammersley sampling, the first variable is sampled according to a $1/N$ scheme, where N is the number of samples. For example, if you are sampling 2 random variables between 0 and 1, with 100 samples, the first variable would have values of $1/100$, $2/100$, etc. This method of sampling causes the matrix $(X'X)$ to become nearly singular, where X is the matrix of sample values. Since one needs to invert the $(X'X)$ matrix to obtain the regression coefficients, we see numerical problems when the condition number of this matrix is very high, for example, around 30. In addition, both Hammersley and Halton sampling can lead to highly correlated values between input variables. For example, the correlation between input dimensions X_1 and X_2 looks very good as shown in Figure 16, but the correlation between X_1 and X_4 shows high correlation as shown in Figure 17:

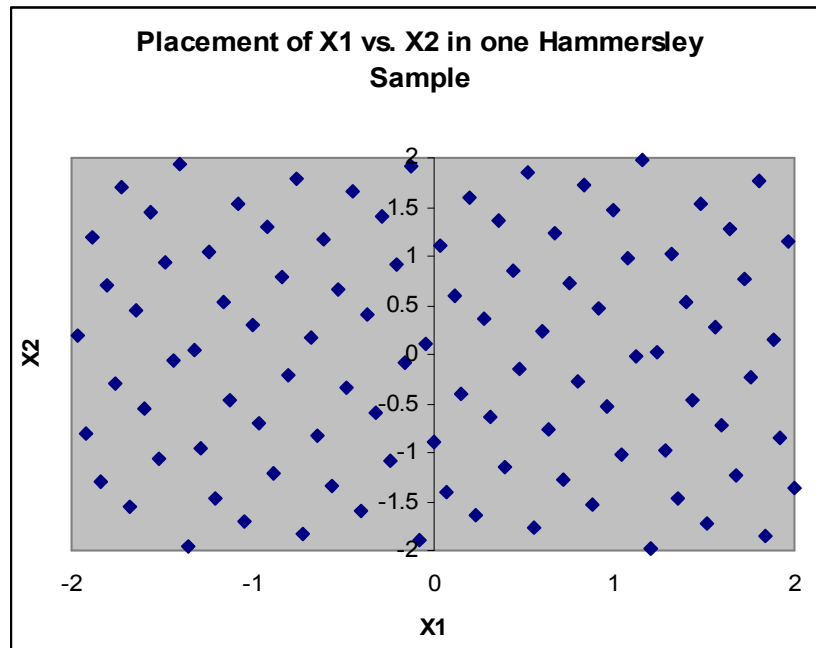


Figure 16. One 100-point sample set showing X_1 vs. X_2 in a Hammersley Sample

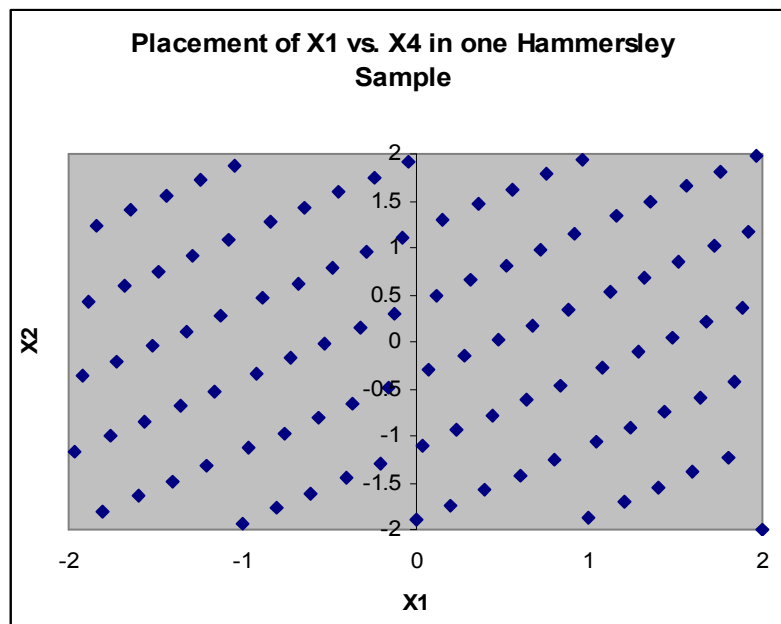


Figure 17. One 100-point sample set showing X_1 vs. X_4 in a Hammersley Sample

We were aware of the possibility of the correlations^{3,17} but we expected it to happen for much larger prime bases than the ones being used to generate these 5-d or 10-d input sets. To avoid correlations amongst inputs in both Halton and Hammersley sequences, it is possible to implement a “fix” which involves skipping values of the sequence when generating data. This is implemented in DAKOTA and is described in the DAKOTA reference manual, however we did not use it in this study.

For the purposes of comparing the other sampling methods, we left the Hammersley sampling out and performed the ANOVA on the regression surfaces for the Paviani function. One can see that CVT does very well with respect to both the RMSE and MAE metrics as shown in Figures 18 and 19.

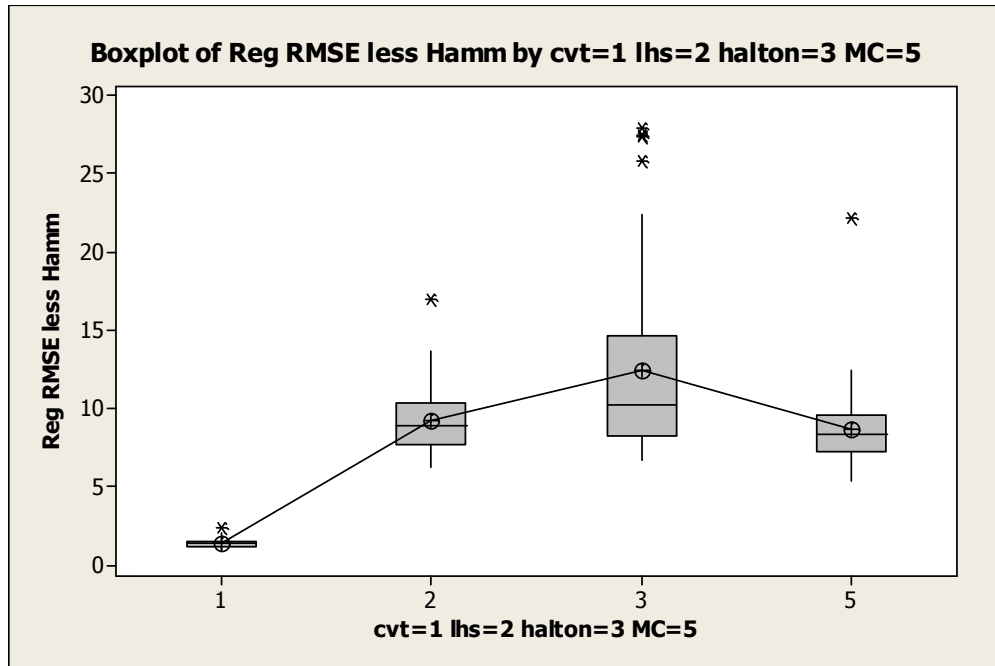


Figure 18. Polynomial Regression RMSE vs. Sample Type omitting Hammersley: Paviani function

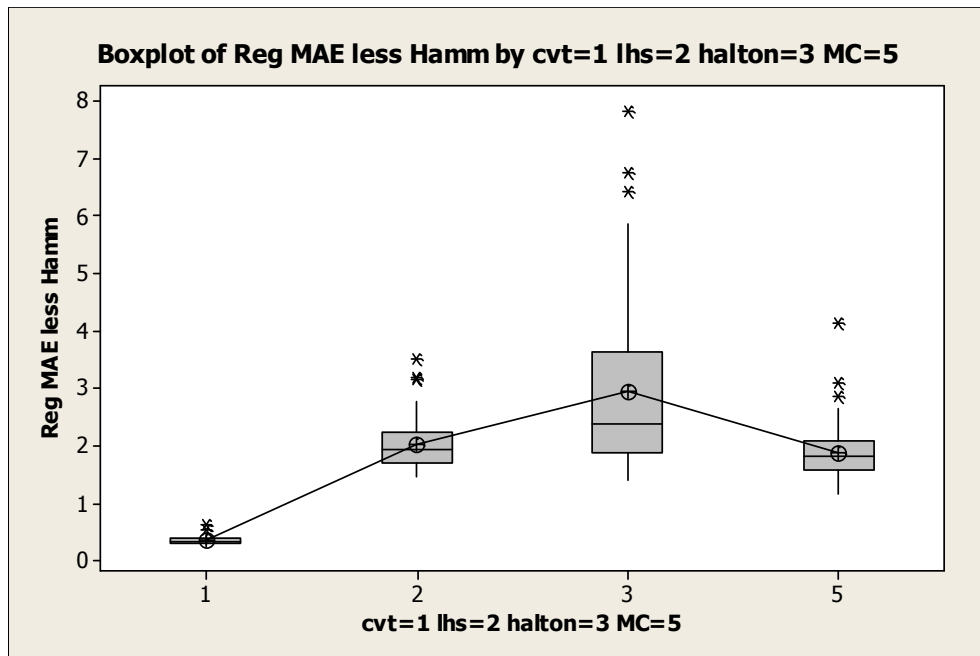


Figure 19. Polynomial Regression MAE vs. Sample Type omitting Hammersley: Paviani function

H. Results of Multivariate Adaptive Regression Spline (MARS) RSA on the Paviani function

Figures 20 and 21 show the results of the RMSE and the MAE for the MARS RSA of the Paviani function. MARS performed very well on this function.

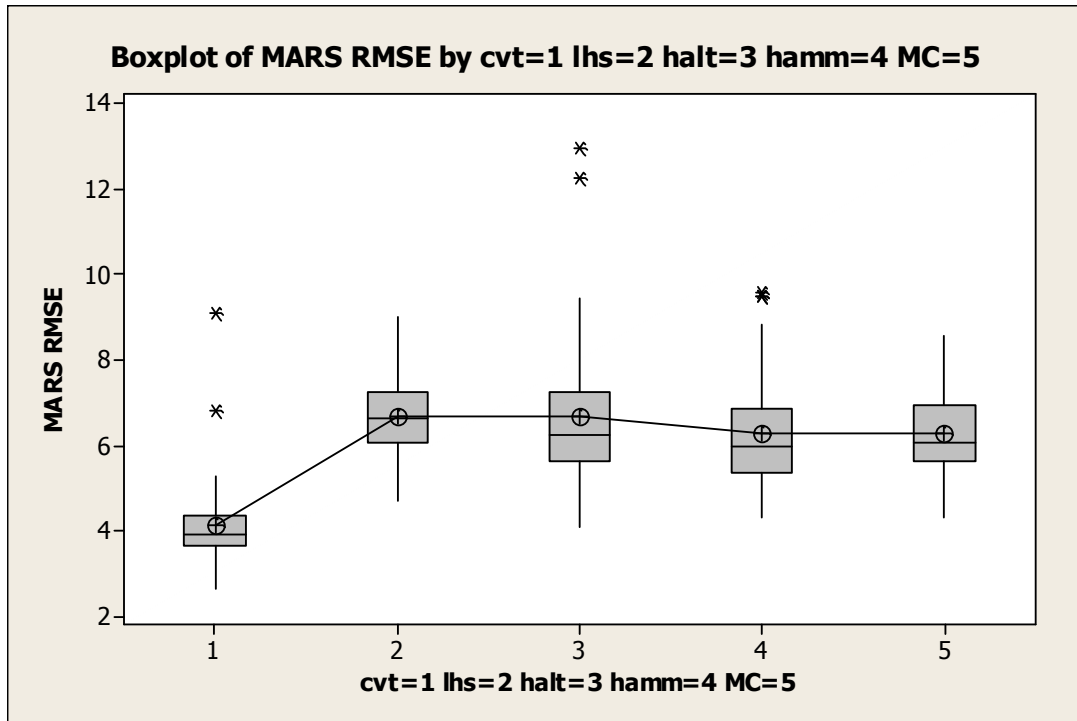


Figure 20. MARS RMSE vs. Sample Type : Paviani function

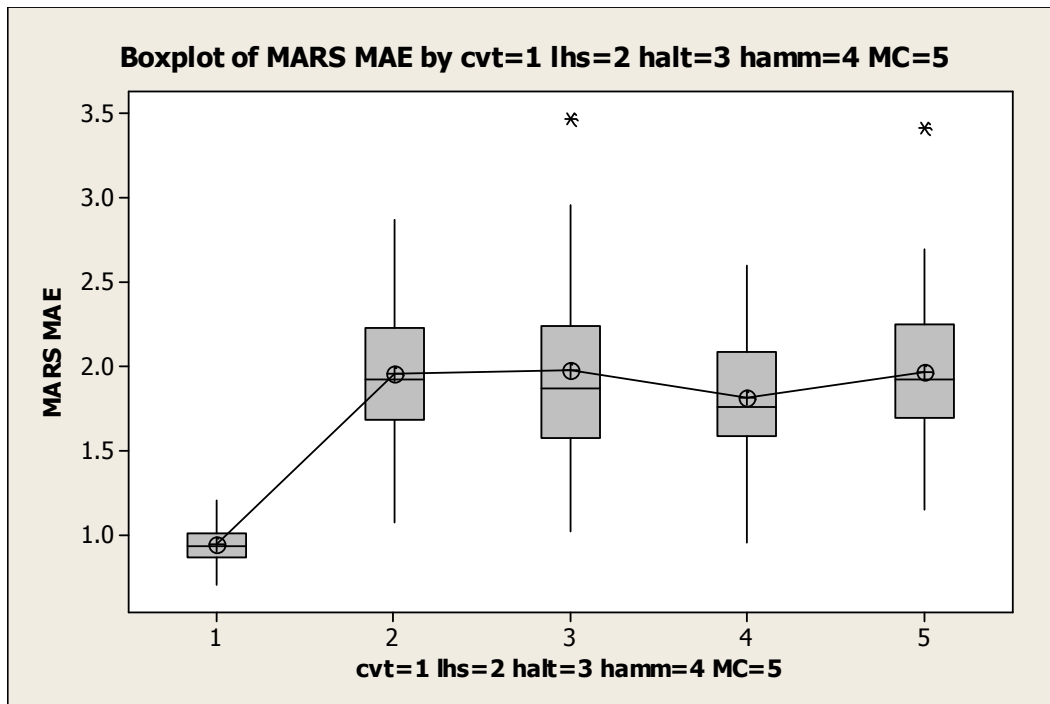


Figure 21. MARS MAE vs. Sample Type : Paviani function

The two-way ANOVA results are shown in Figure 22 for RMSE. Note that we removed Hammersley sampling from this analysis because of the very ill-fitting in the regression case. In general, MARS and kriging were better

response surface approximations than polynomial regression because the Paviani function varies highly in a local region, and the local fitting techniques tend to work better than global methods. CVT again produces a different type of interaction than the other sampling methods. It is interesting to note that regression performs better with CVT than the kriging or MARS.

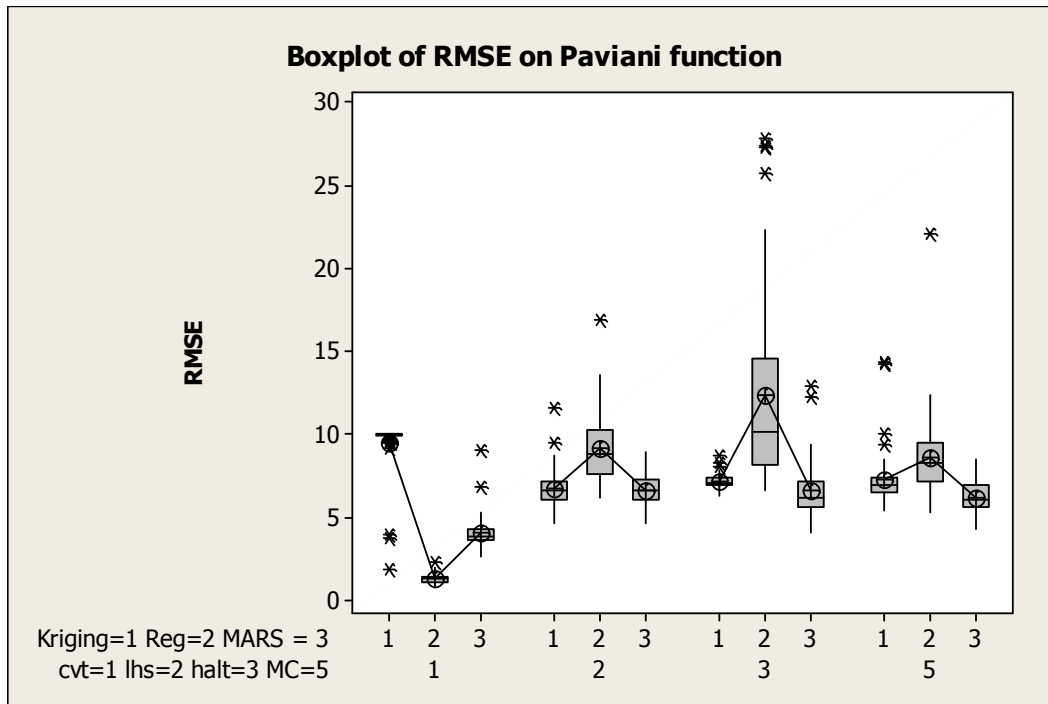


Figure 22. Two-way ANOVA for Paviani Function RMSE

Two-way ANOVA: RMSE versus cvt=1 lhs=2 halt=3 MC=5, Krig=1 Reg=2 MARS = 3

Source	DF	SS	MS	F	P
cvt=1 lhs=2 halt	3	1115.66	371.887	77.47	0.000
Kriging=1 Reg=2	2	470.48	235.242	49.00	0.000
Interaction	6	2653.71	442.285	92.13	0.000
Error	588	2822.78	4.801		
Total	599	7062.64			

cvt=1		
lhs=2		
halt=3		
MC=5	Mean	Individual 95% CIs For Mean Based on Pooled StDev
1	5.00638	(--*)
2	7.55194	(--*)
3	8.77014	(--*)
5	7.39510	(--*)

Kriging=1		
Reg=2		
MARS = 3	Mean	Individual 95% CIs For Mean Based on Pooled StDev
1	7.71176	(---*)
2	7.89772	(---*)
3	5.93319	(---*)

Boxplot of MAE on Paviani Function

MAE

1 2 3 1 2 3 1 2 3 1 2 3

Kriging=1 Reg=2 MARS = 3
cvt=1 lhs=2 halt=3 MC=5

The overall trends are that CVT does best and Halton worst for both MAE and RMSE on the Paviani function. Kriging produced significantly higher (worse) MAE metrics than the MARS or regression, and MARS did much better at producing good RMSE metrics than regression or kriging. Note that for both RMSE and MAE with the Paviani function, there are some interaction effects between the sampling type and response surface type. The interaction was most significant with CVT, since the CVT results per response surface type had different trends than the other sampling types.

V. Summary

Many studies have been done examining the efficacy of sampling methods with respect to various metrics such as uniformity and point placement. However, we have not found many studies which look at the sampling methods and how they perform with various response surface methods. This study has attempted to do that. We examined two problems with 5 input variables (the 5-D version of the Rosenbrock function) and 10 input variables (the Paviani function) respectively. While these may not seem like high dimensional problems, many of the studies in the literature only use 2 or 3 input variables. This work extends our understanding to larger problems. Note that we did not include sample size as a factor under analysis. We fixed the sample size at 100, based on what we have seen done in real applications with approximately the same number of input variables. We did perform extensive tests, in that we generated 50 replicates of the 100 size sample for each sample type, and used those replicates to generate kriging, regression, and MARS response surfaces. The response surface accuracy was then evaluated over a grid with 59049 points to calculate RMSE and MAE, and ANOVA studies were used to look at mean differences in factor levels.

We had hoped to find some clear trends or interactions between sample type and response surface method. We did not. We did see some interesting behavior. CVT produced some unusual results. It was the worst performing sampling method on the Rosenbrock function in terms of RMSE and MAE, but it was the best sampling method on the Paviani function. We feel that CVT warrants further investigation. When we removed CVT from the analysis of the Rosenbrock function, the performance of the other sampling methods (LHS, Halton, Hammersley, and MC) was indistinguishable. This is an important result, saying that the sample type essentially doesn't matter, at least for the Rosenbrock function. Overall, MARS and kriging appeared to produce the best overall fits to the 5-D Rosenbrock function.

With the Paviani function, we found that CVT performed significantly better than the other sampling methods in terms of RMSE and MAE. Halton sampling did not perform well. Hammersley sampling also did not perform well, especially when used with polynomial regression surface approximations. Regression was not able to produce reasonable response surface models with Hammersley samples because of the correlations in inputs. This emphasizes that one needs to be cautious in the choice of sampling method used with response surface type. In the Paviani results, MARS did very well at producing low RMSE and MAE. However, regression did well only on one metric: regression did well minimizing MAE but not RMSE. Kriging did not do particularly well on either. This demonstrates the fact that different response surface approximations have characteristics that will produce different results depending on one's metric of interest: overall scaled accuracy as measured by a mean absolute error, extreme errors which are emphasized in RMSE, overall min or max error, etc.

We did compare our results with those of Simpson, Lin, and Chen²⁰. They also did a comparison of sampling methods and response surface type on two problems, one with 3 input variables and one with 14 input variables. Both problems exhibited nonlinear behavior. Simpson et al. showed some general trends in their paper. For example, the Hammersley sequence gave low results for RMSE, but often was quite high when measuring maximum absolute deviation. Overall, they found that orthogonal arrays and quasi-uniform designs tended to do better than pure random or Latin Hypercube designs. We did not find that in our study. They did mention that Hammersley sampling was not able to be used with regression modeling on their 14 variable problem because of the singular matrix ($X'X$), which agrees with our results. Finally, many of their results had a similar interpretation as ours: certain response surface types worked well on one function but not another, certain sampling types worked well with respect to RMSE but not MAE or vice-versa, etc. Based on our study and Simpson, Lin, and Chen's study, we feel there is no "silver bullet" approach. When using sampling methods to construct response surface approximations, the analyst must be aware of various pitfalls, and ideally use at least two types of sampling and

response surface approximations to understand the characteristics of his or her problem to model it reasonably accurately with a response surface model.

VI. References

1. Booker, Andrew. "Well-conditioned Kriging Models for Optimization of Computer Simulations." *Technical Document Series, M&CT-TECH-002*, Phantom Works, Mathematics and Computing Technology, The Boeing Company, Seattle, WA, 2000.
2. Cressie, N. (1991), *Statistics of Spatial Data*, John Wiley and Sons, New York, NY.
3. Diwekar U. M. and J. R. Kalagnanam (1997). "An Efficient Sampling Technique for Optimization Under Uncertainty", *AIChE Journal*, 43, 440.
4. Du, Q., V. Faber, and M. Gunzburger, 1999. "Centroidal Voronoi Tessellations: Applications and Algorithms," *SIAM Review*, Volume 41, 1999, pages 637-676.
5. Eldred, M.S., Giunta, A.A., van Bloemen Waanders, B.G., Wojtkiewicz, S.F., Jr., Hart, W.E. and Alleva, M.P. (2001), *DAKOTA Users Manual: Version 3.1*, Sandia Technical Report SAND2001-3796, Sandia National Laboratories, Albuquerque, NM. (see: <http://endo.sandia.gov/DAKOTA/software.html>)
6. Friedman, J.H. (1991), "Multivariate Adaptive Regression Splines," *Annals of Statistics*, Vol. 19, No. 1, pp. 1-141.
7. Giunta, A. A., and Watson, L. T., "A Comparison of Approximation Modeling Techniques: Polynomial Versus Interpolating Models," AIAA Paper 98-4758 in *Proceedings of the 7th AIAA/USAF/NASA/ISSMO Symposium on Multidisciplinary Analysis and Optimization*, St. Louis, MO, Sept. 1998, pp. 392-404.
8. Giunta, A. A., Wojtkiewicz, S. F., Jr., and Eldred, M. S. (2003), "Overview of Modern Design of Experiments Methods for Computational Simulations," paper AIAA-2003-0649 in *Proceedings of the 41st AIAA Aerospace Sciences Meeting and Exhibit*, Reno, NV.
9. Halton, J. H. "On the efficiency of certain quasi-random sequences of points in evaluating multi-dimensional integrals, *Numerische Mathematik*, Volume 2 pages 84-90.
10. Halton, J. H. and G. B. Smith, 1964. Algorithm 247: Radical-Inverse Quasi-Random Point Sequence, *Communications of the ACM*, Volume 7, pages 701-702.
11. Helton, J. C. and Davis, F. J. (2001). "Latin Hypercube Sampling and the Propagation of Uncertainty in Analyses of Complex Systems." Technical Report SAND2001-0417, Sandia National Laboratories, Albuquerque, NM.
12. Iman, R.L., and Conover, W.J. (1982a). "Sensitivity Analysis Techniques: Self-Teaching Curriculum," Nuclear Regulatory Commission Report, NUREG/CR-2350, Technical Report SAND81-1978, Sandia National Laboratories, Albuquerque, NM.
13. Iman, R.L., and Conover, W.J. (1982b). "A Distribution-Free Approach to Inducing Rank Correlation Among Input Variables," *Communications in Statistics*, B11(3), 311-334.
14. Kocis, L. and W. Whiten, 1997. "Computational Investigations of Low-Discrepancy Sequences," *ACM Transactions on Mathematical Software*, Volume 23, Number 2, 1997, pages 266-294.
15. McKay, M.D., Beckman, R.J., and Conover, W.J. (1979), "A Comparison of Three Methods for Selecting Values of Input Variables in the Analysis of Output from a Computer Code," *Technometrics*, Vol. 21, No. 2, pp. 239-245.
16. Metropolis N. and Ulam, S., "The Monte Carlo Method," *Journal of the American Statistical Association*, Vol. 44, No. 247, 1949, pp. 335-341.
17. Robinson, D.G. and C. Atcitty, 1999. "Comparison of Quasi- and Pseudo-Monte Carlo Sampling for Reliability and Uncertainty Analysis." *Proceedings of the AIAA Probabilistic Methods Conference*, St. Louis MO, AIAA99-1589.
18. Romero, V.J., Burkardt, J.V., Gunzburger, M.D., and J.S. Peterson (2003). "Initial Evaluation of Pure and "Latinized" Centroidal Voronoi Tessellation for Non-Uniform Statistical Sampling." SAMO Conference Paper, 2003.
19. Romero, V. J., Swiler, L. P., and Giunta, A. A., "Construction of Response Surfaces Based on Progressive-Lattice-Sampling Experimental," *Structural Safety*, Vol. 26, No. 2, 2004, pp. 201-219.
20. Simpson, T., Dennis, L., and Chen, W. (2002). "Sampling Strategies for Computer Experiments: Design and Analysis" *Journal of International Journal of Reliability and Application*, 2(3), 209-240.
21. Swiler, L. P. and Wyss, G. D., "A User's Guide to Sandia's Latin Hypercube Sampling Software: LHS UNIX Library/Standalone Version," Sandia Technical Report SAND2004-2439, Sandia National Laboratories, Albuquerque, NM, 2004.